

Scaling and Taylor's law for heavy tailed observations

with J. Cohen and R. Davis

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- ▶ **Taylor's law:** for some positive constants a , b :

$$S_n^2 \approx a(\bar{X}_n)^b \quad \text{for moderately large } n.$$

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- ▶ We focus on Taylor's law and related scaling resulting from heavy tails.

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- ▶ X is in $\text{RV}(\alpha)$ means $\bar{F}(x) = x^{-\alpha} L(x)$, L slowly varying.

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- ▶ S_α : a totally skewed to the right α -stable random variable.

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- ▶ $(S_\alpha, S_{\alpha/2})$ marginally stable, jointly infinitely divisible.

$$S_\alpha = \begin{cases} \sum_{j=1}^{\infty} \left(\Gamma_j^{-1/\alpha} - E[\Gamma_j^{-1/\alpha} \mathbf{1}(\Gamma_j^{-1/\alpha} \leq 1)] \right) & \text{if } 1 \leq \alpha < 2 \\ \sum_{j=1}^{\infty} \Gamma_j^{-1/\alpha} & \text{if } 0 < \alpha < 1 \end{cases} .$$

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► $\Gamma_1, \Gamma_2, \dots$: standard Poisson arrivals.

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- ▶ $a_n = n^{1/\alpha} L_1(n)$, L_1 slowly varying.

$$(L_1(n))^{b-2} W_n = \frac{a_n^{-2} \sum_{i=1}^n (X_i - \bar{X}_n)^2}{(a_n^{-1} \sum_{i=1}^n X_i)^b}$$

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an approximate Taylor's law.

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 - ▶ Z_0, Z_1, Z_2, \dots iid standard normals, $0 \leq \rho < 1$, $0 < \alpha < 1$.

$$X_i = \frac{1}{|\rho^{1/2}Z_0 + (1 - \rho)^{1/2}Z_i|^{1/\alpha}}, \quad i = 1, 2, \dots$$

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$$\frac{S_n^2}{\bar{X}_n^b} \Rightarrow \frac{S_{\alpha/2}}{S_\alpha^b} \left(\frac{2}{\pi}\right)^{\alpha/(2(1-\alpha))} \left[(1-\rho) \exp\left\{ \frac{\rho}{1-\rho} Z_0^2 \right\} \right]^{1/(2(1-\alpha))}.$$

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$$\hat{\kappa}_n = \frac{n^{-1} \sum_{i=1}^n (X_i - \bar{X}_n)^4}{\left(n^{-1} \sum_{i=1}^n (X_i - \bar{X}_n)^2 \right)^2}.$$

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- ▶ If X_1, X_2, \dots iid, $EX^4 < \infty$, then $\hat{\kappa}_n \rightarrow \kappa$ a.s.

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- ▶ If variances finite: $\rho \rightarrow \rho$ a.s.
- ▶ What happens when variances are infinite?

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$(S_{\alpha/2,12}, S_{\alpha/2,1}, S_{\alpha/2,2})$: a 3-dimensional $\alpha/2$ -stable random vector

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- ▶ $(S_{1,j}, S_{2,j}), j = 1, 2, \dots$: iid random vectors on $\{s_1^2 + s_2^2 = 1, s_1, s_2 \geq 0\}$, independent of the Poisson arrivals $\Gamma_j, j \geq 1$.

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normalized to be a probability measure.

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equivalently, X and Y are asymptotically independent.

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- ▶ In some cases normalized ρ_n has a nondegenerate limit.
- ▶ **Example:** (Z_1, Z_2) is bivariate normal, zero means, unit variances, correlation $\rho \in (-1, 1)$.

Set $X = |Z_1|^{-c}$, $Y = |Z_2|^{-c}$, some $c > 1$.

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$$R_1 \stackrel{d}{=} \frac{\sum_{j=1}^{\infty} \Gamma_j^{-c}}{\sqrt{\sum_{j=1}^{\infty} \Gamma_j^{-2c}}}.$$

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- ▶ $((X_i, Y_i), i \geq 1)$ iid copies of (X, Y) .
- ▶ The least squares estimator of η :

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- ▶ Then

$$\left(\frac{n}{\log n}\right)^{1/\alpha} (\eta - \hat{\eta}_n) \Rightarrow C^{1/\alpha} \frac{\sum_{j=1}^{\infty} \Gamma_j^{-1/\alpha}}{\sum_{j=1}^{\infty} \Gamma_j^{-2/\alpha}}$$

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a much faster rate of convergence.

Taylor's law, with correlations, $\alpha = 0.1$

Regularly varying tails, $b=2.1111$

