

Parameter estimation of discretely observed interacting particle systems

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joint work with
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Plan of this talk

1. Model
2. Problem
3. Assumptions
4. Examples
5. Main results & proofs
6. Open problems

1. Model

Applications in

- ▶ mathematical biology (large societies of animals, neural networks)
- ▶ social sciences (multiagents' opinion, consensus)
- ▶ finance (systemic risk)
- ▶ mean-field games (a large number of players)...

An interacting particle system:

$$(X_t^{1,N}, \dots, X_t^{N,N})_{t \in [0, T]},$$

which satisfies

$$\begin{cases} dX_t^{i,N} = b(X_t^{i,N}, \mu_t^N, \theta_1) dt + a(X_t^{i,N}, \mu_t^N, \theta_2) dW_t^i, & i = 1, \dots, N, t \in [0, T], \\ \mathcal{L}(X_0^{1,N}, \dots, X_0^{N,N}) = \mu_0 \times \dots \times \mu_0, \end{cases}$$

where

- ▶ $W_t^i := (W_t^i)_{t \in [0, T]}$, $i = 1, \dots, N$, are independent standard Brownian motions,
- ▶ $X_0^{1,N}, \dots, X_0^{N,N}$ are independent of W^1, \dots, W^N ,
- ▶ $\mu_t^N := N^{-1} \sum_{i=1}^N \delta_{X_t^{i,N}}$,
- ▶ $b : \mathbb{R} \times \mathcal{P}_2(\mathbb{R}) \times \Theta_1 \rightarrow \mathbb{R}$, $a : \mathbb{R} \times \mathcal{P}_2(\mathbb{R}) \times \Theta_2 \rightarrow \mathbb{R}$, where $\mathcal{P}_2(\mathbb{R})$ is the set of all probability measures on \mathbb{R} with finite 2nd moments endowed with the Wasserstein 2-metric W_2^* , $\Theta_j \subset \mathbb{R}^{p_j}$ is compact, convex for some $p_j \in \mathbb{N}$, $j = 1, 2$,
- ▶ $\theta_j \in \Theta_j^\circ$, $j = 1, 2$, are parameters of interest.

* $W_2(\mu, \nu) := \inf_{\pi \in \Pi(\mu, \nu)} (\int_{\mathbb{R} \times \mathbb{R}} |x - y|^2 \pi(dx, dy))^{1/2}$, where $\Pi(\mu, \nu)$ denotes the set of all probability measures on $\mathbb{R} \times \mathbb{R}$ with marginals μ, ν

The associated McKean-Vlasov SDE (introduced by McKean 1966):

$$\begin{cases} d\bar{X}_t = b(\bar{X}_t, \bar{\mu}_t, \theta_1)dt + a(\bar{X}_t, \bar{\mu}_t, \theta_2)dW_t, & t \in [0, T], \\ \bar{\mu}_0 = \mu_0, \end{cases} \quad (1)$$

where

- ▶ W is a standard Brownian motion,
- ▶ \bar{X}_0 is independent of W ,
- ▶ $\bar{\mu}_t := \mathcal{L}(\bar{X}_t)$.

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The **propagation of chaos property**: if \bar{X}^i satisfies (1) with $W = W^i$, $\bar{X}_0 = X_0^{i,N}$, $i = 1, \dots, N$, under A1, A2 (to be stated later), then

$$\mathbb{E}[|X_t^{i,N} - \bar{X}_t^i|^2] \leq CN^{-1/2}, \quad i = 1, \dots, N, t \in [0, T].$$

This and other probabilistic properties: Sznitman 1984, 1991, Méléard 1996, Fernandez, Méléard 1997, Tanaka, Hitsuda 1981, Malrieu 2001, Bolley, Guillin, Villani 2007, Cattiaux, Guillin, Malrieu 2008, ...

2. Problem

Statistical inference:

- ▶ parametric: Kasonga 1990, Löcherbach 2002, Bishwal 2011, Giesecke, Schenkler, Sirignano 2020, Genon-Catalot, Larédo 2021a, 2021b, 2022+, Chen 2021, Sharrock, Kantas, Parpas, Pavliotis 2021+, Della Maestra, Hoffmann 2022+
- ▶ nonparametric: Della Maestra, Hoffmann 2022, Belomestny, Pilipauskaitė, Podolskij 2022+, Belomestny, Podolskij, Zhou 2022+

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Problem of this talk: **estimation of the parameter $\theta := (\theta_1, \theta_2)$** from observations

$$X_{j\Delta_n}^{i,N}, \quad i = 1, \dots, N, \quad j = 1, \dots, n,$$

with $\Delta_n := T/n$, where T is fixed and $N, n \rightarrow \infty$.

We define an estimator of $\theta = \theta_0$ as

$$\hat{\theta}_n^N \in \arg \min_{\theta \in \Theta} S_n^N(\theta),$$

where

$$S_n^N(\theta) := \sum_{0 < i \leq N} \sum_{0 \leq j < n} \left\{ \frac{(X_{(j+1)\Delta_n}^{i,N} - X_{j\Delta_n}^{i,N} - b(X_{j\Delta_n}^{i,N}, \mu_{j\Delta_n}^N, \theta_1)\Delta_n)^2}{a^2(X_{j\Delta_n}^{i,N}, \mu_{j\Delta_n}^N, \theta_2)\Delta_n} + \log a^2(X_{j\Delta_n}^{i,N}, \mu_{j\Delta_n}^N, \theta_2) \right\}$$

uses a Gaussian quasi-likelihood which originates from the Euler scheme:

$$X_{(j+1)\Delta_n}^{i,N} \approx X_{j\Delta_n}^{i,N} + b(X_{j\Delta_n}^{i,N}, \mu_{j\Delta_n}^N, \theta_1)\Delta_n + a(X_{j\Delta_n}^{i,N}, \mu_{j\Delta_n}^N, \theta_2)(W_{(j+1)\Delta_n}^i - W_{j\Delta_n}^i).$$

The Euler scheme contrast in case of the classical SDE: Florens-Zmirou 1989, Yoshida 1992, Kessler 1997

3. Assumptions

A1 μ_0 has finite k -th moment for all $k \in \mathbb{N}$.

A2 For all θ there exists a $C > 0$ such that all $(x, \mu), (y, \nu)$,

$$|b(x, \mu, \theta_1) - b(y, \nu, \theta_1)| + |a(x, \mu, \theta_2) - a(y, \nu, \theta_2)| \leq C(|x - y| + W_2(\mu, \nu)).$$

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A3 There exists a $C > 0$ such that $a^2(x, \mu, \theta_2) > C$ for all (x, μ, θ_2) .

A4 (i) $b(x, \mu, \cdot) \in C^3(\Theta_1)$, $a(x, \mu, \cdot) \in C^3(\Theta_2)$ for all (x, μ) and their partial derivatives up to order 3 have polynomial growth in (x, μ) uniformly in θ .

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(ii) For all θ there exist $C > 0$, $k, l = 0, 1, \dots$ such that for all $r_1 + r_2 = 1, 2$, $h_1, h_2 = 1, \dots, p_1$, $\tilde{h}_1, \tilde{h}_2 = 1, \dots, p_2$ and $(x, \mu), (y, \nu)$,

$$\begin{aligned} & \left| \partial_{\theta_1, h_1}^{r_1} \partial_{\theta_1, h_2}^{r_2} b(x, \mu, \theta_1) - \partial_{\theta_1, h_1}^{r_1} \partial_{\theta_1, h_2}^{r_2} b(y, \nu, \theta_1) \right| \\ & + \left| \partial_{\theta_2, \tilde{h}_1}^{r_1} \partial_{\theta_2, \tilde{h}_2}^{r_2} a(x, \mu, \theta_2) - \partial_{\theta_2, \tilde{h}_1}^{r_1} \partial_{\theta_2, \tilde{h}_2}^{r_2} a(y, \nu, \theta_2) \right| \\ & \leq C(|x - y| + W_2(\mu, \nu))(1 + |x|^k + |y|^k + W_2^l(\mu, \delta_0) + W_2^l(\nu, \delta_0)). \end{aligned}$$

A5 For all $\varepsilon > 0$,

$$\inf_{\theta \in \Theta: \|\theta_1 - \theta_{0,1}\| \geq \varepsilon} I(\theta) > 0 \quad \text{and} \quad \inf_{\theta_2 \in \Theta_2: \|\theta_2 - \theta_{0,2}\| \geq \varepsilon} J(\theta_2) > J(\theta_{0,2}),$$

where $I : \Theta \rightarrow \mathbb{R}$, $J : \Theta_2 \rightarrow \mathbb{R}$ are defined as

$$I(\theta) := \int_0^T \int_{\mathbb{R}} \frac{(b(x, \bar{\mu}_t, \theta_1) - b(x, \bar{\mu}_t, \theta_{0,1}))^2}{a^2(x, \bar{\mu}_t, \theta_2)} \bar{\mu}_t(dx) dt,$$
$$J(\theta_2) := \int_0^T \int_{\mathbb{R}} \left(\frac{a^2(x, \bar{\mu}_t, \theta_{0,2})}{a^2(x, \bar{\mu}_t, \theta_2)} + \log a^2(x, \bar{\mu}_t, \theta_2) \right) \bar{\mu}_t(dx) dt.$$

A6 The main-diagonal block elements of $\Sigma(\theta_0) := \text{diag}(\Sigma^{(1)}(\theta_0), \Sigma^{(2)}(\theta_0))$ are $p_j \times p_j$ invertible $\Sigma^{(j)}(\theta_0) = (\Sigma_{kl}^{(j)}(\theta_0))$, $j = 1, 2$, where

$$\Sigma_{kl}^{(1)}(\theta_0) := 2 \int_0^T \int_{\mathbb{R}} \frac{\partial_{\theta_{1,k}} b(x, \bar{\mu}_t, \theta_{0,1}) \partial_{\theta_{1,l}} b(x, \bar{\mu}_t, \theta_{0,1})}{a^2(x, \bar{\mu}_t, \theta_{0,2})} \bar{\mu}_t(dx) dt,$$

$$\Sigma_{kl}^{(2)}(\theta_0) := \int_0^T \int_{\mathbb{R}} \frac{\partial_{\theta_{2,k}} a^2(x, \bar{\mu}_t, \theta_{0,2}) \partial_{\theta_{2,l}} a^2(x, \bar{\mu}_t, \theta_{0,2})}{a^4(x, \bar{\mu}_t, \theta_{0,2})} \bar{\mu}_t(dx) dt.$$

A7 There are some $\tilde{a}, K \in C^2(\mathbb{R}^2)$ whose partial derivatives of order 1, 2 have polynomial growth such that for all (x, μ) ,

$$a(x, \mu, \theta_{0,2}) := \tilde{a}\left(x, \int_{\mathbb{R}} K(x, y) \mu(dy)\right).$$

4. Examples

- (i) The Kuramoto model: N oscillators defined by N angles $X_t^{i,N}$ (defined modulo 2π) evolving over $[0, T]$ according to

$$dX_t^{i,N} = -\frac{\theta_1}{N} \sum_{j=1}^N \sin(X_t^{i,N} - X_t^{j,N}) dt + \theta_2 dW_t^i.$$

Here

$$b(x, \mu, \theta_1) = -\theta_1 \int_{\mathbb{R}} \sin(x - y) \mu(dy), \quad a(x, \mu, \theta_2) = \theta_2.$$

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- (ii) A model for opinion dynamics:

$$dX_t^{i,N} = -\frac{1}{N} \sum_{j=1}^N \varphi(|X_t^{i,N} - X_t^{j,N}|, \theta_1) (X_t^{i,N} - X_t^{j,N}) dt + \theta_2 dW_t^i,$$

where the influence function $\varphi(x, \theta_1)$, which acts on the “difference of opinions” between agents, is an infinitely differentiable approximation of $\theta_{1,1} \mathbf{1}_{[0, \theta_{1,2}]}(x)$, $x \in \mathbb{R}$.

(iii) The interacting particle system

$$dX_t^{i,N} = \left(\theta_{1,1} + \frac{\theta_{1,2}}{N} \sum_{j=1}^N X_t^{j,N} - \theta_{1,3} X_t^{i,N} \right) dt + \theta_2 \sqrt{1 + (X_t^{i,N})^2} dW_t^i$$

for $\theta_{1,2} = 0$ reduces to N independent samples of a special case of the Pearson diffusion, which has applications in finance.

(iv) The mean field limit of

$$dX_t^{i,N} = \left(\theta_{1,1} + \frac{\theta_{1,2}}{N} \sum_{j=1}^N X_t^{j,N} - \theta_{1,3} X_t^{i,N} \right) dt + \left(\theta_{2,1} + \theta_{2,2} \sqrt{\frac{1}{N} \sum_{j=1}^N (X_t^{j,N})^2} \right) dW_t^i$$

is a time-inhomogeneous Ornstein-Uhlenbeck process. See Kasonga 1990 for $\theta_{1,1} = \theta_{2,2} = 0$.

5. Main results & proofs

Theorem (Consistency)

Assume A1-A5 without A4(ii). Then

$$\hat{\theta}_n^N \xrightarrow{\mathbb{P}} \theta_0 \quad \text{as } N, n \rightarrow \infty.$$

Theorem (Asymptotic normality)

Assume A1-A7. If $N\Delta_n \rightarrow 0$ then

$$\left(\sqrt{N}(\hat{\theta}_{n,1}^N - \theta_{0,1}), \sqrt{N/\Delta_n}(\hat{\theta}_{n,2}^N - \theta_{0,2}) \right) \xrightarrow{\mathcal{L}} \mathcal{N}\left(0, 2(\Sigma(\theta_0))^{-1}\right) \quad \text{as } N, n \rightarrow \infty.$$

Technical results: Assume A1-A2.

- Then for all $k \in \mathbb{N}$, $N \in \mathbb{N}$, $i = 1, \dots, N$, $s, t \in [0, T]$ it holds

$$\sup_{t \in [0, T]} \mathbb{E}[|X_t^{i, N}|^k] \leq C, \quad \mathbb{E}[|X_t^{i, N} - X_s^{i, N}|^k] \leq C|t - s|^{k/2}.$$

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- Moreover, assume that $f : \mathbb{R} \times \mathcal{P}_l(\mathbb{R}) \rightarrow \mathbb{R}$ satisfies for some $C > 0$, $k, l = 0, 1, \dots$ and all $(x, \mu), (y, \nu)$,

$$\begin{aligned} & |f(x, \mu) - f(y, \nu)| \\ & \leq C(|x - y| + W_2(\mu, \nu))(1 + |x|^k + |y|^k + W_l^l(\mu, \delta_0) + W_l^l(\nu, \delta_0)) \end{aligned}$$

and $(x, t) \mapsto f(x, \bar{\mu}_t)$ is in $L^1(\mathbb{R} \times [0, T], \bar{\mu}_t(dx)dt)$. Then as $N, n \rightarrow \infty$,

$$\frac{\Delta_n}{N} \sum_{0 < i \leq N} \sum_{0 \leq j < n} f(X_{j\Delta_n}^{i, N}, \mu_{j\Delta_n}^N) \xrightarrow{\mathbb{P}} \int_0^T \int_{\mathbb{R}} f(x, \bar{\mu}_t) \bar{\mu}_t(dx) dt.$$

Proof of consistency:

- ▶ Assume A1-A4(i). Then

$$\begin{aligned} \sup_{\theta \in \Theta} \left| \frac{\Delta_n}{N} S_n^N(\theta) - J(\theta_2) \right| &= o_{\mathbb{P}}(1), \\ \sup_{\theta \in \Theta} \left| \frac{1}{N} (S_n^N(\theta) - S_n^N(\theta_{0,1}, \theta_2)) - I(\theta) \right| &= o_{\mathbb{P}}(1). \end{aligned} \quad (2)$$

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- ▶ A5 implies that for every $\varepsilon > 0$ there exists $\eta > 0$ such that

$$\{\|\hat{\theta}_{n,1}^N - \theta_{0,1}\| \geq \varepsilon\} \subseteq \{I(\hat{\theta}_n^N) > \eta\},$$

where the definition of $\hat{\theta}_n^N$ and (2) imply

$$I(\hat{\theta}_n^N) \leq I(\hat{\theta}_n^N) - \frac{1}{N} (S_n^N(\hat{\theta}_n^N) - S_n^N(\theta_{0,1}, \hat{\theta}_n^N)) = o_{\mathbb{P}}(1).$$

The proof is similar in the case of $\hat{\theta}_{n,2}^N$.

Proof of asymptotic normality: By mean value theorem and definition of $\hat{\theta}_n^N$,

$$-\nabla_{\theta} S_n^N(\theta_0) = (\hat{\theta}_n^N - \theta_0) \int_0^1 \nabla_{\theta}^2 S_n^N(\theta_0 + s(\hat{\theta}_n^N - \theta_0)) ds.$$

Furthermore, we introduce

$$M_n^N := \text{diag} \left(N^{-\frac{1}{2}} \mathbf{1}_{p_1}, \left(\frac{\Delta_n}{N} \right)^{\frac{1}{2}} \mathbf{1}_{p_2} \right),$$
$$\Sigma_n^N(\theta) := M_n^N \nabla_{\theta}^2 S_n^N(\theta) M_n^N.$$

Steps:

- ▶ If $N\Delta_n \rightarrow 0$ then

$$\nabla_{\theta} S_n^N(\theta_0) M_n^N = \sum_{0 < j \leq n} \xi_{n,j} \xrightarrow{\mathcal{L}} \mathcal{N}(0, 2\Sigma(\theta_0))$$

follows from

$$\sum_{0 < j \leq n} \mathbb{E}_{(j-1)\Delta_n} [\xi_{n,j}] \xrightarrow{\mathbb{P}} 0$$

and a central limit theorem for martingale difference triangular arrays (Theorem 3.2 in Hall, Heyde 1980):

$$\sum_{0 < j \leq n} \mathbb{E}_{(j-1)\Delta_n} [\xi_{n,j} (\xi_{n,j})^{\top}] \xrightarrow{\mathbb{P}} 2\Sigma(\theta_0), \quad \sum_{0 < j \leq n} \mathbb{E}_{(j-1)\Delta_n} [|\xi_{n,j,h}|^4] \xrightarrow{\mathbb{P}} 0,$$

where $h = 1, \dots, p_1 + p_2$ and $\mathbb{E}_t[\cdot]$ denotes the conditional expectation with respect to $\sigma(\{X_0^{i,N}, (W_s^i)_{s \in [0,t]}, i = 1, \dots, N\})$.

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where $h = 1, \dots, p_1 + p_2$ and $\mathbb{E}_t[\cdot]$ denotes the conditional expectation with respect to $\sigma(\{X_0^{i,N}, (W_s^i)_{s \in [0,t]}, i = 1, \dots, N\})$.

- ▶ $\Sigma_n^N(\theta_0) \xrightarrow{\mathbb{P}} \Sigma(\theta_0)$,
- ▶ $\sup_{s \in [0,1]} \|\Sigma_n^N(\theta_0 + s(\hat{\theta}_n^N - \theta_0)) - \Sigma_n^N(\theta_0)\| \xrightarrow{\mathbb{P}} 0$.

6. Open problems

- ▶ Efficiency
- ▶ To improve $N\Delta_n \rightarrow 0$
- ▶ $X_t^{i,N} \in \mathbb{R}^d$
- ▶ Presence of jumps
- ▶ Non or semi-parametric estimation

References

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Merci pour votre attention