

Bootstrap for integer-valued GARCH processes

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Presentation plan

1 Integer-valued GARCH

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- 1 Integer-valued GARCH
- 2 Properties of an INGARCH process

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Integer-valued GARCH

Classical GARCH (Bollerslev, 1986):

$$\begin{aligned}X_t &= \sigma_t \varepsilon_t, \\ \sigma_t^2 &= \omega + \sum_{i=1}^p \alpha_i X_{t-i}^2 + \sum_{j=1}^q \beta_j \sigma_{t-j}^2,\end{aligned}$$

where $(\varepsilon_t)_{t \in \mathbb{Z}}$ i.i.d., $E\varepsilon_t = 0$, $E[\varepsilon_t^2] = 1$.

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$$\leadsto \text{Var}(X_t^2 \mid X_{t-1}, \sigma_{t-1}, X_{t-2}, \sigma_{t-2}, \dots) = \omega + \sum_{i=1}^p \alpha_i X_{t-i}^2 + \sum_{j=1}^q \beta_j \sigma_{t-j}^2.$$

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Integer-valued counterpart: Poisson-INGARCH

$$\begin{aligned}X_t \mid \mathcal{F}_{t-1} &\sim \text{Pois}(\lambda_t), \\ \lambda_t &= \omega + \sum_{i=1}^p \alpha_i X_{t-i} + \sum_{j=1}^q \beta_j \lambda_{t-j},\end{aligned}$$

where $\mathcal{F}_s = \sigma(X_s, \lambda_s, X_{s-1}, \lambda_{s-1}, \dots)$

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Our framework: Nonlinear Poisson-INGARCH

$$\begin{aligned} X_t \mid \mathcal{F}_{t-1} &\sim \text{Pois}(\lambda_t), \\ \lambda_t &= f_{\theta_0}(X_{t-1}, \dots, X_{t-p}, \lambda_{t-1}, \dots, \lambda_{t-p}) \end{aligned}$$

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Condition 1 f_{θ_0} “contractive”
 $\exists c_1, \dots, c_p, d_1, \dots, d_q \geq 0, \sum_{i=1}^p c_i + \sum_{j=1}^q d_j < 1$:

$$\begin{aligned} &|f_{\theta_0}(x_1, \dots, x_p, \lambda_1, \dots, \lambda_q) - f_{\theta_0}(x'_1, \dots, x'_p, \lambda'_1, \dots, \lambda'_q)| \\ &\leq \sum_{i=1}^p c_i |x_i - x'_i| + \sum_{j=1}^q d_j |\lambda_j - \lambda'_j| \end{aligned}$$

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Properties of the INGARCH process

Define $Z_t := (X_t, \dots, X_{t-p+1}, \lambda_t, \dots, \lambda_{t-q+1})$.

Then $\mathbf{Z} = (Z_t)_{t \in \mathbb{Z}}$ is a time-homogeneous Markov process with state space $S := \mathbb{N}_0^p \times [0, \infty)^q$.

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Implications:

- \mathbf{Z} has a unique stationary distribution
- $(X_t)_{t \in \mathbb{Z}}$ is absolutely regular (β -mixing)

Coupling

Metric on S :

$$\begin{aligned}\Delta_{\gamma,\delta}((x_1, \dots, x_p, \lambda_1, \dots, \lambda_q), (x'_1, \dots, x'_p, \lambda'_1, \dots, \lambda'_q)) \\ := \sum_{i=1}^p \gamma_i |x_i - x'_i| + \sum_{j=1}^q \delta_j |\lambda_j - \lambda'_j|\end{aligned}$$

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Proposition 1 (Markov kernel is contractive)

Suppose that Condition 1 is fulfilled. Then, for an appropriate choice of $\gamma_1, \dots, \gamma_p, \delta_1, \dots, \delta_q > 0$, $\kappa < 1$, there exist

$$Z \sim P_{\theta_0}^{Z_t|Z_{t-1}=z} \quad \text{and} \quad Z' \sim P_{\theta_0}^{Z_t|Z_{t-1}=z'}$$

on a probability space $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{P})$ such that

$$\tilde{E} \Delta_{\gamma,\delta}(Z, Z') \leq \kappa \Delta_{\gamma,\delta}(z, z'). \quad (2.1)$$

Proof of Proposition 1 For $z = (x_1, \dots, x_p, \lambda_1, \dots, \lambda_q)$,
 $z' = (x'_1, \dots, x'_p, \lambda'_1, \dots, \lambda'_q)$, define

$$Z := (X, x_1, \dots, x_{p-1}, \lambda, \lambda_1, \dots, \lambda_{q-1}),$$

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where, according to the model equation,

$$\lambda = f_{\theta_0}(x_1, \dots, x_p, \lambda_1, \dots, \lambda_q), \quad \lambda' = f_{\theta_0}(x'_1, \dots, x'_p, \lambda'_1, \dots, \lambda'_q).$$

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$$\text{(Condition 1)} \quad \rightsquigarrow \quad |\lambda - \lambda'| \leq \sum_{i=1}^p c_i |x_i - x'_i| + \sum_{j=1}^q d_j |\lambda_j - \lambda'_j|$$

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To generate $X \sim \text{Pois}(\lambda)$ and $X' \sim \text{Pois}(\lambda')$, take a Poisson process with unit intensity: $(N(u))_{u \geq 0}$ and define

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Then, there exist $\gamma_1, \dots, \gamma_p, \delta_1, \dots, \delta_q > 0$, $\kappa < 1$ such that

$$\tilde{\Delta}_{\gamma, \delta}(Z, Z') \leq \dots \leq \kappa \Delta_{\gamma, \delta}(z, z').$$

Another representation of the contraction property

Kantorovich (Wasserstein L^1) distance:

$$\mathcal{K}(Q, Q') := \inf_{Z \sim Q, Z' \sim Q'} \tilde{E} \Delta_{\gamma, \delta}(Z, Z'),$$

where Z and Z' are random variables with respective distributions Q and Q' , both defined on a common probability space $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{P})$.

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Proposition 2 Suppose that Condition 1 is fulfilled. Then, for arbitrary distributions Q, Q' on S ,

$$\mathcal{K}(Q\pi^Z, Q'\pi^Z) \leq \kappa \mathcal{K}(Q, Q'),$$

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Proof Follows from (2.1). □

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Corollary 1

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Proof

- $Q \mapsto Q\pi^Z$ contractive
- $\mathcal{P} := \{Q: Q \text{ probab. distr.}, \int \|x\| dQ(x) < \infty\}$ complete

Banach fixed point theorem: π^Z admits a unique fixed point Q_0 , i.e. $Q_0\pi^Z = Q_0$. □

Absolute regularity of the count process

Corollary 2

Suppose that Condition 1 is fulfilled. Then $(X_t)_{t \in \mathbb{Z}}$ is absolutely regular (β -mixing),

$$\beta_X(n) \leq C \rho^n \quad \forall n \in \mathbb{N},$$

for some $C < \infty$, $\rho < 1$.

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Proof

$$\beta_X(n)$$

$$= E \left[\sup_C \left| P((X_n, X_{n+1}, \dots) \in C \mid X_0, X_{-1}, \dots) - P((X_n, X_{n+1}, \dots) \in C) \right| \right]$$

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$$\begin{aligned} \beta_X(n) &= E \left[\sup_C \left| P((X_n, X_{n+1}, \dots) \in C \mid X_0, X_{-1}, \dots) - P((X_n, X_{n+1}, \dots) \in C) \right| \right] \\ &\leq E \left[\sup_C \left| P((X_n, X_{n+1}, \dots) \in C \mid Z_0, Z_{-1}, \dots) - P((X_n, X_{n+1}, \dots) \in C) \right| \right] \\ &= E \left[\sup_C \left| P((X_n, X_{n+1}, \dots) \in C \mid Z_0) - P((X_n, X_{n+1}, \dots) \in C) \right| \right]. \end{aligned}$$

Proof of Corollary 2 (contd.)

If $(\tilde{Z}_t)_{t \in \mathbb{N}_0}$ and $(\tilde{Z}'_t)_{t \in \mathbb{N}_0}$ are two versions of the process $(Z_t)_{t \in \mathbb{N}_0}$ such that \tilde{Z}_0 and \tilde{Z}'_0 are independent, then

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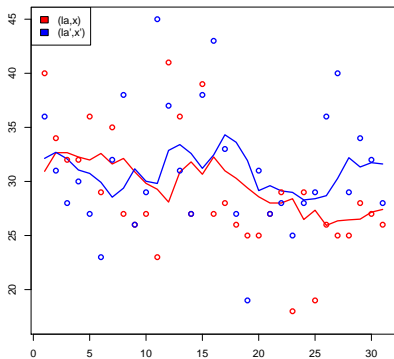
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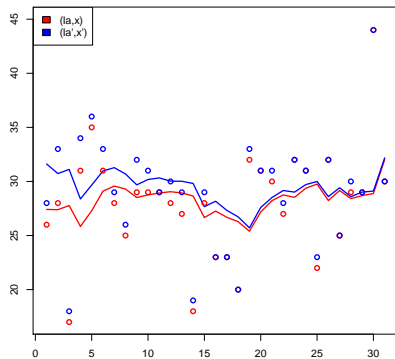
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□

INGARCH(1,1): $(I_{a,x})$ and $(I_{a',x'})$ independently run



INGARCH(1,1): $(I_{a,x})$ and $(I_{a',x'})$ coupled



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- Sometimes knowledge of the properties of $(X_t)_{t \in \mathbb{N}}$ required:
 - confidence intervals/sets (width)
 - hypothesis tests (critical value)
- “Bootstrap”:
Given x_1, \dots, x_n , construct an artificial process $(X_t^*)_{t \in \mathbb{N}}$ which (hopefully) mimics the behavior of $(X_t)_{t \in \mathbb{N}}$

Bootstrap (contd.)

In our case:

- parametric model:

$$\begin{aligned}X_t \mid \mathcal{F}_{t-1} &\sim \text{Pois}(\lambda_t), \\ \lambda_t &= f_{\theta_0}(X_{t-1}, \dots, X_{t-p}, \lambda_{t-1}, \dots, \lambda_{t-q})\end{aligned}$$

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(e.g. by conditional maximum likelihood)
- construct bootstrap process:
 - choose pre-sample values $X_0^*, \dots, X_{1-p}^*, \lambda_0^*, \dots, \lambda_{1-q}^*$
 - for $t = 1, \dots, n$:

$$\lambda_t^* = f_{\widehat{\theta}_n}(X_{t-1}^*, \dots, X_{t-p}^*, \lambda_{t-1}^*, \dots, \lambda_{t-q}^*)$$

$$X_t^* \sim \text{Pois}(\lambda_t^*)$$

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- parametric model:

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$$\lambda_t = f_{\theta_0}(X_{t-1}, \dots, X_{t-p}, \lambda_{t-1}, \dots, \lambda_{t-q})$$

- estimate θ_0 be $\hat{\theta}_n = \hat{\theta}_n(X_1, \dots, X_n)$
(e.g. by conditional maximum likelihood)
- construct bootstrap process:
 - choose pre-sample values $X_0^*, \dots, X_{1-p}^*, \lambda_0^*, \dots, \lambda_{1-q}^*$
 - for $t = 1, \dots, n$:

$$\lambda_t^* = f_{\hat{\theta}_n}(X_{t-1}^*, \dots, X_{t-p}^*, \lambda_{t-1}^*, \dots, \lambda_{t-q}^*)$$
$$X_t^* \sim \text{Pois}(\lambda_t^*)$$

- hope for the best:

$$P_{\theta_0}^{X_1^*, \dots, X_n^* \mid X_1, \dots, X_n} \quad \approx \quad P_{\theta_0}^{X_1, \dots, X_n}$$

Bootstrap consistency

Usual approach:

- $S_n = S_n(X_1, \dots, X_n; \theta_0)$ statistic of interest,
e.g. $S_n = \sqrt{n}(\bar{X}_n - EX_1)$, $\bar{X}_n = n^{-1} \sum_{t=1}^n X_t$.
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More general approach:

Construct, on $(\tilde{\Omega}, \tilde{\mathcal{F}}, \tilde{P})$, respective versions $(\tilde{Z}_t)_{t=1, \dots, n}$ and $(\tilde{Z}_t^*)_{t=1, \dots, n}$ of $(Z_t)_{t=1, \dots, n}$ and $(Z_t^*)_{t=1, \dots, n}$ such that

$$\tilde{E} \Delta_{\gamma, \delta}(\tilde{Z}_t, \tilde{Z}_t^*) \quad \text{“small”}$$

Conditions for bootstrap consistency

Condition 1'

$\exists c_1, \dots, c_p, d_1, \dots, d_q \geq 0, \sum_{i=1}^p c_i + \sum_{j=1}^q d_j < 1:$

$$\begin{aligned} & \left| f_{\theta}(x_1, \dots, x_p, \lambda_1, \dots, \lambda_p) - f_{\theta}(x'_1, \dots, x'_q, \lambda'_1, \dots, \lambda'_q) \right| \\ & \leq \sum_{i=1}^p c_i |x_i - x'_i| + \sum_{j=1}^q d_j |\lambda_j - \lambda'_j| \quad \forall \theta \in \Theta_0, \end{aligned}$$

where $\Theta_0 = \{\theta \in \Theta: \|\theta - \theta_0\| \leq \delta\}, \delta > 0.$

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Condition 3

$$\begin{aligned} & |f_{\theta}(x_1, \dots, x_p, \lambda_1, \dots, \lambda_q) - f_{\theta_0}(x_1, \dots, x_p, \lambda_1, \dots, \lambda_q)| \\ & \leq M \|\theta - \theta_0\| \left(\sum_{i=1}^p x_i + \sum_{j=1}^q \lambda_j \right), \end{aligned}$$

for all $(x_1, \dots, x_p, \lambda_1, \dots, \lambda_q) \in S$ and all $\theta \in \Theta_0$.

A characterization of bootstrap consistency

Theorem 1

Let Conditions 1', 2 and 3 be fulfilled. Then

$$\begin{aligned} & \tilde{E} \Delta_{\gamma, \delta}(\tilde{Z}_t^*, \tilde{Z}_t) \\ & \leq \frac{ME \|Z_0\| (\gamma_1 + \delta_1)}{1 - \kappa} \|\hat{\theta}_n - \theta_0\| + \kappa^t \tilde{E} \Delta_{\gamma, \delta}(\tilde{Z}_0^*, \tilde{Z}_0) \end{aligned}$$

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$$\tilde{Z}_0 = (\tilde{X}_0, \dots, \tilde{X}_{1-p}, \tilde{\lambda}_0, \dots, \tilde{\lambda}_{1-q}), \quad \tilde{Z}_0^* = (\tilde{X}_0^*, \dots, \tilde{X}_{1-p}^*, \tilde{\lambda}_0^*, \dots, \tilde{\lambda}_{1-q}^*)$$

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For $t \geq 1$,

$$\begin{aligned} \tilde{\lambda}_t &= f_{\theta_0}(\tilde{Z}_{t-1}), & \tilde{\lambda}_t^* &= f_{\hat{\theta}_n}(\tilde{Z}_{t-1}^*) \\ \tilde{X}_t &= N_t(\tilde{\lambda}_t), & \tilde{X}_t^* &= N_t(\tilde{\lambda}_t^*), \end{aligned}$$

where $(N_1(u))_{u \geq 0}$, $(N_2(u))_{u \geq 0}$, ... independent Poisson processes

Properties of the coupled process

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Let Conditions 1', 2 and 3 be fulfilled. If $\widehat{\theta}_n \in \Theta_0$, then

- (i) $((\widetilde{Z}_t, \widetilde{Z}_t^*))_{t \in \mathbb{Z}}$ has a unique stationary distribution.

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Draw on

$$\widetilde{E}[\Delta_{\gamma, \delta}(\widetilde{Z}_t, \widetilde{Z}'_t) + \Delta_{\gamma, \delta}(\widetilde{Z}_t^*, \widetilde{Z}'_t{}^*)] \leq \kappa \widetilde{E}[\Delta_{\gamma, \delta}(\widetilde{Z}_{t-1}, \widetilde{Z}'_{t-1}) + \Delta_{\gamma, \delta}(\widetilde{Z}_{t-1}^*, \widetilde{Z}'_{t-1}{}^*)].$$

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- generalized means

- autocovariances

- degenerate von Mises and U -statistics

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Details: Neumann, M. H. (2021). Bootstrap for integer-valued generalized GARCH(p,q) processes. *Statistica Neerlandica* **75** (3), 343–363.