

*Inheritance of strong mixing and weak
dependence under renewal sampling*

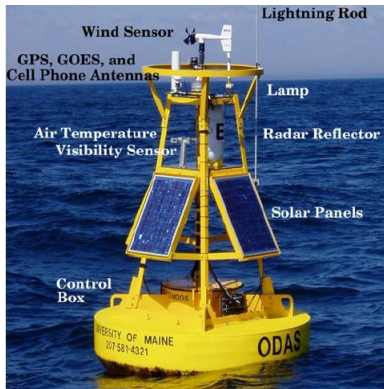
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based on a joint work with D. Brandes and R. Stelzer

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EcoDep 2021 Conference



The sentinel of the sea



Point Reference Data: sea surface temperature

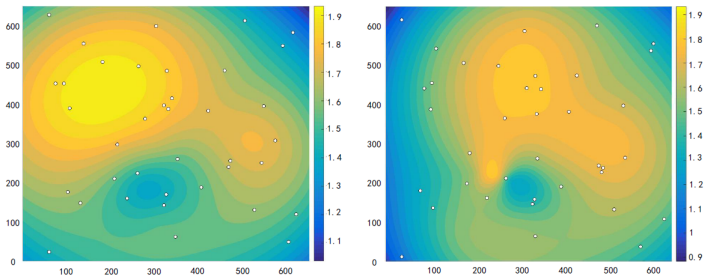


Figure: Spatio-Temporal reference point data in two time stamps. The white circles illustrates the locations of the sensors recording the temperature field values. The set $\{(t_s, x_s, Z_s) \text{ for } s = 1, \dots, N\}$ is called point reference data set. Source: Wang et al. (2019), Deep learning for spatio-temporal data mining: a survey.

Following the reading of one sensor

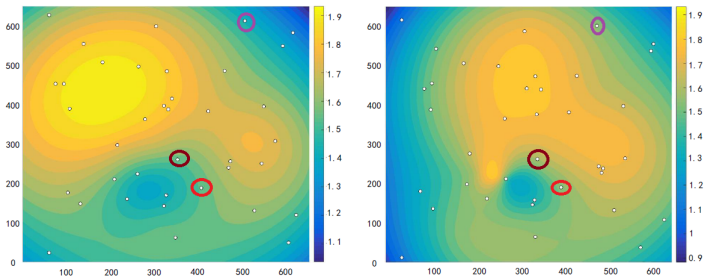


Figure: Trajectories data

Spatio-temporal trajectories data

Non-equidistant
(random) data
in time and/or
space which are
serially correlated

How to
model?

Time series

- Mobile, networked sensors can also be carried by people, (e.g., smartphones) or animals (e.g, animal tracking), enabling the monitoring of heart rate, body temperature, among other information.

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We study renewal sampling of $(X_t)_{t \in \mathcal{I}}$

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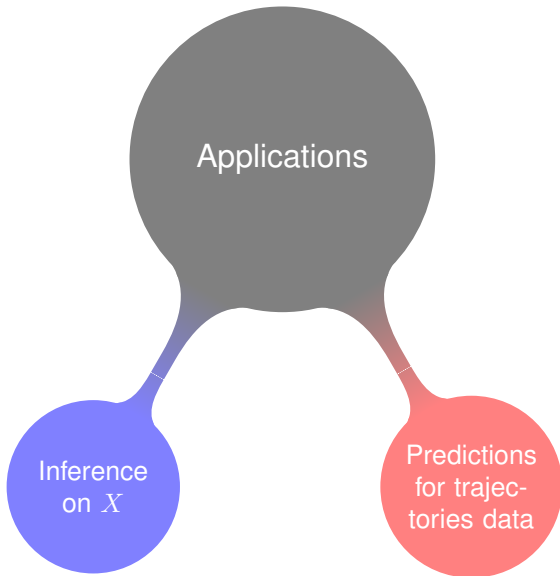
Results

We show that if

● X is strictly-stationary, $\eta, \lambda, \kappa, \zeta, \theta$ -weakly dependent, see Dedecker et al. (2008);



● A renewal sampling of X is inheriting the dependence structure of X .



Renewal Sampling

Let $\mathcal{I} \subseteq \mathbb{R}^m$ and $\tau = (\tau_i)_{i \in \mathbb{Z} \setminus \{0\}}$ be an \mathcal{I} -valued sequence of non-negative (component-wise) i.i.d. random vectors with distribution function μ such that $\mu\{0\} < 1$. For $i \in \mathbb{Z}$, we define an \mathcal{I} -valued stochastic process $(T_i)_{i \in \mathbb{Z}}$ as

$$T_0 := 0 \quad \text{and} \quad T_i := \begin{cases} \sum_{j=1}^i \tau_j, & i \in \mathbb{N}, \\ -1 - \sum_{j=i}^{-1} \tau_j, & -i \in \mathbb{N}. \end{cases}$$

The sequence $(T_i)_{i \in \mathbb{Z}}$ is called a renewal sampling sequence.



Independence assumption

Independence assumption

- TS*: Observation times depend on the measuring instrument (typically sensors), i.e., on a random source independent of the process X , as observed by Bardet and Bertrand (2010).
- ST*: The sampling in space-time depends on the source of randomness proper of the instrument used to record them.

Renewal sampled process

Let $X = (X_t)_{t \in \mathcal{I}}$ and let $(T_i)_{i \in \mathbb{Z}}$ be a renewal sampling sequence independent of X . We define the sequence $Y = (Y_i)_{i \in \mathbb{Z}}$ as the stochastic process with values in \mathbb{R}^{d+1} given by

$$Y_i = \begin{pmatrix} X_{T_i} \\ \tau_i \end{pmatrix}.$$

We call X the underlying process and Y the renewal sampled process.

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Remark: This modeling is designed to work when the sampling scheme is not known, i.e., it is not designed by an experimenter but just observed from the data.

Definition of Ψ -weak dependence

For any positive integer u, v , and functions F and G being bounded Lipschitz or bounded measurable functions, weakly dependent processes (or random fields) satisfy covariance inequalities of the following type:

$$\begin{aligned} |Cov(F(X_{i_1}, \dots, X_{i_u}), G(X_{j_1}, \dots, X_{j_v}))| & \quad (1) \\ & \leq c \Psi(\|F\|_\infty, \|G\|_\infty, Lip(F), Lip(G), u, v) \epsilon(r), \end{aligned}$$

where



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where

- the sequence of coefficients $\epsilon = (\epsilon(r))_{r \in \mathbb{R}^+}$ converges to zero as $r \rightarrow \infty$,
- c is a constant independent of r and $dist(\{i_1, \dots, i_u\}, \{j_1, \dots, j_v\}) \geq r$,
- the function $\Psi(\cdot)$ has different shapes depending on the functional spaces where F and G are defined, and the dependence notion under analysis.



Theorem (Brandes, C., Stelzer)

Let $Y = (Y_i)_{i \in \mathbb{Z}}$ be a \mathbb{R}^{d+1} -valued process with $X = (X_t)_{t \in \mathcal{I}}$ being strictly-stationary and Ψ -weakly dependent with coefficients $\epsilon = (\epsilon(r))_{r \in \mathbb{R}^+}$. Then, it exists a sequence $(\mathcal{E}(n))_{n \in \mathbb{N}^*}$ satisfying

$$\begin{aligned} |Cov(\tilde{F}(Y_{i_1}, \dots, Y_{i_u}), \tilde{G}(Y_{j_1}, \dots, Y_{j_v}))| \\ \leq C \Psi(\|\tilde{F}\|_\infty, \|\tilde{G}\|_\infty, Lip(\tilde{F}), Lip(\tilde{G}), u, v) \mathcal{E}(n) \end{aligned}$$

where C is a constant independent of n , $dist(\{i_1, \dots, i_u\}, \{j_1, \dots, j_v\}) \geq n$, and \tilde{F}, \tilde{G} are either bounded Lipschitz or bounded measurable function.

Moreover,

$$\mathcal{E}(n) = \int_{\mathcal{I}} \epsilon(\|r\|) \mu^{*n}(dr), \quad (2)$$

with μ^{*n} the n -fold convolution of μ .



Weak dependent coefficients

If X is strictly stationary and η -weakly dependent

$$\begin{aligned}
 & |Cov(\tilde{F}(Y_{i_1}, \dots, Y_{i_u}), \tilde{G}(Y_{j_1}, \dots, Y_{j_v}))| \\
 & \leq C u Lip(\tilde{F}) \|\tilde{G}\|_\infty + v Lip(\tilde{G}) \|\tilde{F}\|_\infty \mathcal{E}(n)
 \end{aligned}$$

Proposition (Brandes, C. and Stelzer)

- For F and G bounded measurable functions, and

$$\Psi(\|F\|_\infty, \|G\|_\infty, Lip(F), Lip(G), u, v) = \|F\|_\infty \|G\|_\infty$$

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Remark: we perform an alternative proof to the one of Charlot and Rachdi (2007).



θ -lex weak dependence, (C., Stelzer, and Ströh (2021))

- Lexicographic order on \mathbb{R}^m : for distinct elements $y = (y_1, \dots, y_m) \in \mathbb{R}^m$ and $z = (z_1, \dots, z_m) \in \mathbb{R}^m$ we say $y <_{lex} z$ if and only if $y_1 < z_1$ or $y_p < z_p$ for some $p \in \{2, \dots, m\}$ and $y_q = z_q$ for $q = 1, \dots, p - 1$.

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- Let F be a bounded and G bounded Lipschitz functions, and $I_u = \{i_1, \dots, i_u\} \subset \mathbb{R}^m$, and $j \in \mathbb{R}^m$ be such that $i_s <_{lex} j$ for all $s = 1, \dots, u$, and $dist(I_u, j) \geq r$. Then,

$$\Psi(\|F\|_\infty, \|G\|_\infty, Lip(F), Lip(G), u, 1) = \|F\|_\infty Lip(G),$$

and ϵ corresponds to the **θ -lex-coefficients**.

Corollary (Brandes, C., Stelzer)

Let X be a strictly stationary and θ -lex weakly dependent random field defined on \mathbb{R}^m , and $\tau = (\tau_i)_{i \in \mathbb{Z} \setminus \{0\}}$ be an \mathbb{R}^m -valued sequence of non-negative (component-wise) i.i.d. random vector with distribution function μ . Then Y is a strictly stationary process, and there exists a sequence \mathcal{E} such that

$$|Cov(\tilde{F}(Y_{i_1}, \dots, Y_{i_u}), \tilde{G}(Y_j))| \leq C \|F\|_\infty Lip(G) \mathcal{E}(n)$$

where C is a constant independent of n , and \mathcal{E} are defined in (2).

Corollary (Brandes, C. and Stelzer)

If the coefficients $(\mathcal{E}(n))_{n \in \mathbb{N}}$ are finite, and converge to zero as n goes to infinity, then Y is Ψ -weakly dependent with coefficients \mathcal{E} .

Ψ -weakly dependent renewal sampled processes

Proposition (Brandes, C. and Stelzer)

- **Exponential Decay:** Let us assume that $\epsilon(r) \leq Ce^{-\gamma r}$ for $\gamma > 0$ and denote the Laplace transform of the distribution function μ of the inter-arrival time τ by

$$\mathcal{L}_\mu(t) = \int_{\mathbb{R}_+} e^{-tr} \mu(dr), \quad t \in \mathbb{R}_+.$$

Then, the process Y admits coefficients

$$\mathcal{E}(n) \leq C \left(\frac{1}{\mathcal{L}_\mu(\gamma)} \right)^{-n}$$

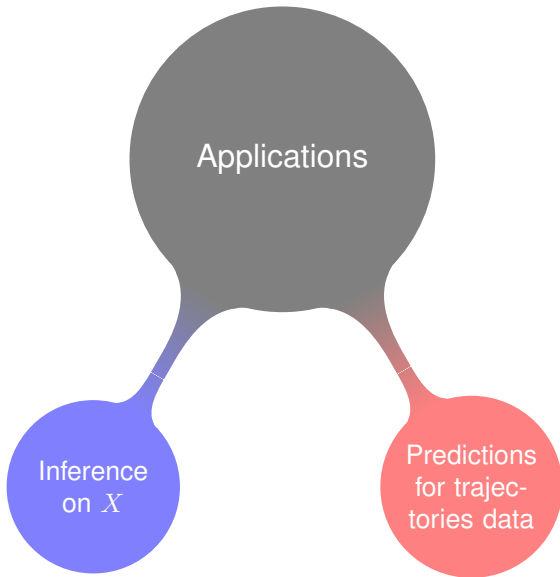
which converge to zero as n goes to infinity.

Ψ -weakly dependent renewal sampled processes

Proposition (Brandes, C. and Stelzer)

- **Power decay:** Let us assume that $\epsilon(r) \leq Cr^{-\gamma}$ for $\gamma > 0$. Let $a > 0$ be a point in the support of the distribution function μ of the inter-arrival time τ such that $\mu([0, a)) > 0$, and set $p = \mu([a, \infty])$. Then, the process Y admits coefficients

$$\mathcal{E}(n) \leq C(nap)^{-\gamma} \quad \text{as } n \rightarrow \infty.$$



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- central limit theorems exist for a sample $(X_n)_{n \in \mathbb{Z}}$.
- Knowing the dependence structure of Y enable us to check if the above condition hold for the sample $(Y_n)_{n \in \mathbb{Z}}$.

Trajectory data

- Let $\{(t_1^i, x_1^i), \dots, (t_p^i, x_p^i) : t_j^i \in \mathbb{R} \text{ and } x_j^i \in \mathbb{R}^2 \text{ for } i = 1, \dots, N \text{ and } j = 1, \dots, p\}$ represents a set of m different trajectories observed in p space-time points.
- $S = \{(X_1, Y_1), (X_2, Y_2), \dots, (X_N, Y_N)\}$ a generic *training set* where each *example* (X_i, Y_i) is determined by an input-output pair that we assume generated by a spatio-temporal random field $Z = (Z_t(x))_{(t,x) \in \mathbb{R} \times \mathbb{R}^2}$ such that

$$X_i = T_p(x_p^i), \quad \text{and} \quad Y_i = Z_{t_p^i}(x_p^i),$$

where $T_p(x_p^i) = (Z_{t_{p-1}^i}(x_{p-1}^i), \dots, Z_{t_1^i}(x_1^i))$ represents the past of the observation $Z_{t_p^i}(x_p^i)$ along the trajectory i .

Generalization bounds for trajectory data

Future Work:

- Linear predictors.
- Squared and absolute loss functions.

Thank you





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